

## Statistical Approaches in Analysis of Variance: from Random Arrangements to Latin Square Experimental Design

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### Abstract

**Background:** The choices of experimental design as well as of statistical analysis are of huge importance in field experiments. These are necessary to be correctly in order to obtain the best possible precision of the results. The random arrangements, randomized blocks and Latin square designs were reviewed and analyzed from the statistical perspective of error analysis. **Material and Method:** Random arrangements, randomized block and Latin squares experimental designs were used as field experiments. A series of previously published data were analyzed. An algorithm for errors analysis was developed and applied on the experimental data. **Results:** The analysis revealed that the errors classification in random arrangements is:  $\text{Error}(\text{Treatment}) < \text{Error}(\text{Total}) < \text{Error}(\text{Experiment})$ . The errors classification in randomized blocks revealed to be:  $\text{Error}(\text{Treatment}) < \text{Error}(\text{Total}) < \text{Error}(\text{Experiment}) < \text{Error}(\text{Block})$ . The obtained errors classification in Latin square was as follows:  $\text{Error}(\text{Experiment}) < \text{Error}(\text{Treatment}) < \text{Error}(\text{Total}) < \text{Error}(\text{Column}) < \text{Error}(\text{Row})$ . **Conclusions:** The Latin square design proved to have the smallest experimental errors

compared to randomized arrangement and randomized block design. The classification of errors proved to be similar in randomized arrangements and randomized block design.

### **Keywords**

Design of experiment; Random arrangements; Randomized blocks; Latin squares; Analysis of errors.

### **Introduction**

The experimental design could be consider the golden state of any experiment due to its aims to ensure that the experiment is able to detect the treatment effects that are of interest by using the available resources to obtain the best possible precision. The choice of design as well as the choice of statistical analysis can make a huge difference. The methods of experimental design are obviously at least as important as method of data analysis.

First systematic experiment (a clinical trial) was developed and applied by James Lind in 1747 to treat the scurvy (deficiency of vitamin C) [1, 2]. Sir Ronald A. Fisher was the first statistician that applied a formal experimental mathematical model [3].

The peculiarity of field experiments lies in the fact, verified in all careful uniformity trials, that the area of ground chosen for the experimental plots may be assumed to be markedly heterogeneous, in that its fertility varies in a systematic, and often a complicated manner from point to point [4, 5]. This fact gives a risk to be in error when conclusions were drawn from a field experiment. Consequently, in order to be able to obtain statistically significant results from observations, the experiment should be designed in such manner in which the experiment settings (excepting the location for an observable) to be present more than once in the field (the confidence of the observation result increasing with the number of the locations of the observation). The direct way of overcoming this difficulty is to arrange the plots wholly at random.

The paper aims to present and analyze the error from the statistical point of view from random arrangements, randomized blocks and Latin square experimental designs.

## Material and Method

First requirement of a well-designed experiment is that the experiment should yield a comparison of different manures, treatments, varieties and of testing the significance of such differences as are observed [4].

### *Randomized Design*

The randomization is used in field experiments in order to avoid systematic, selection, accidental biases and to avoid the cheating by the experimenter.

The experimental data obtained by Mercer and Hall [6] in a uniformity trial on 20 plots and 5 treatments expressed as weights of mangold root were subject of random arrangements analysis. The following order was used in the experiment: B1 (3504), C1 (3430), A1 (3376), C2 (3334), E1 (3253), E2 (3314), E3 (3287), A2 (3361), D1 (3404), A3 (3366), B2 (3416), C3 (3291), B3 (3244), D2 (3210), D3 (3168), B4 (3195), A4 (3330), D4 (3118), C4 (3029), E4 (3085).

### *Randomized Block Design*

The accuracy of the observation could be increased by using blocks (adding restrictions on the orders of strips arrangements). Thus, the analysis of randomized blocks was also performed. The variance is split three parts in randomized blocks design [7]:

- Local differences between blocks;
- Differences due to treatment;
- Experimental errors.

Thus, the estimate of experimental error becomes an unbiased estimate of the actual error in the differences due to treatment. A random arrangement (or systematic order) of in 4 blocks by imposing the condition that each treatment shall occur once in each block could lead to the following design:

- Block 1: 3314 (A1) - 3416 (E1) - 3404 (C1) - 3029 (D1) - 3504 (B1)
- Block 2: 3334 (C2) - 3207 (B2) - 3195 (E2) - 3287 (D2) - 3361 (A2)
- Block 3: 3118 (A3) - 3085 (D3) - 3244 (E3) - 3167 (B3) - 3366 (C3)
- Block 4: 3291 (C4) - 3330 (E4) - 3253 (B4) - 3376 (A4) - 3430 (D4)

The number of runs needed in complete randomized block designs [8] is directly related with the number of levels of factors implied in analysis. The general formula of number of runs is:

$$n_{\text{run}} = L_1 \times L_{i+1} \times \dots \times L_k \tag{Eq.1.}$$

where  $n_{\text{run}}$  = number of runs;  $L_i$  = number of levels of factor  $i$  ( $1 \leq i \leq k$ ,  $k$  = number of factors). For example when we have 2 factors (treatment factor with five levels and blocking factor with 4 levels) with one replication per plot it will be needed a number of runs equal to  $L_1 \times L_2 = 5 \times 4 = 20$  runs.

The following assumptions were done in error analysis in randomized blocks design ( $H_0$ : the variation within block is the same with the variation within each treatment vs.  $H_a$ : the variation within block is not the same with the variation within each treatment or  $H_0$ : the means of treatments are equal vs  $H_a$ : at least two means differ):

- the blocks are considered random;
- the measurement values follows a normal distribution with the same variance;
- the measurement errors are independent of the block effect; the block effects follows an identical normal distribution with a mean 0.

### Latin Squares

Latin square design and the related Graeco-Latin square and Hyper-Graeco-Latin square designs are a special type of comparative design used to control the variation related to both rows and columns in the field experiment [9]. These experimental designs are used when:

- The analysis of several nuisance factors is desired and it is not proper to combine these factors into a single factor.
- The research resources allows a relative small number of runs:

Design	Number of factors	Number of runs
3×3 Latin square	3	9
3×3 Graeco-Latin Square	4	9
4×4 Latin square*	3	16
4×4 Graeco-Latin Square	4	16
4×4 Hyper-Graeco-Latin Square	5	16

\* equivalent to a  $4^{3-1}$  fractional factorial design [10]

The design is limited to the experiments in which the number of levels of each blocking variable is equal with the number of levels of the treatment factor. Moreover, it could be applied under the assumption that there are no interactions between the blocking variables or between the treatment variable and the blocking variables. Note that the rows and columns are orthogonal to treatments in the Latin square design.

The general model of the response for a Latin square design is presented in Eq. 2.

$$Y_{ijk} = \mu + R_i + C_j + T_k + RE \quad \text{Eq.2.}$$

where  $Y_{ijk}$  = any observation for which  $X_1 = i$ ,  $X_2 = j$  ( $X_1, X_2$  = blocking factors),  $X_3$  = treatment factor;  $\mu$  = general location parameter,  $R_i$  = effect for block  $i$ ;  $C_j$  = effect for block  $j$ ;  $T_k$  = effect for treatment;  $RE$  = random error (measure of the sum of variation between plots or units receiving same treatments).

The experimental data obtained by Mercer and Hall [6] in a block of 25 plots arranged in 5 rows and 5 columns, to be used for testing 5 treatments (each treatment occurs once in each row) was subject to Latin square analysis of errors (Table 1.).

Table 1. Experimental data: Latin square experimental design [6]

	Block 1	Block 2	Block 3	Block 4	Block 5
D	376	E 371	C 355	B 356	A 335
B	316	D 338	E 336	A 356	C 332
C	326	A 326	B 335	D 343	E 330
E	317	B 343	A 330	C 327	D 336
A	321	C 332	D 317	E 318	B 306

## Results and Discussion

### *Randomized Design*

If twenty observations are made at random for a single factor of primary interest for five values of the factor, then a random succession of the factor values must be assured such that exactly four repetitions of every value should be found in the succession.

The summary of the analysis of variance on the investigated sample is presented in Table 2.

The significance of numbers presented in Table 2 is as follows:

1. Sum the observations on every treatment: give the total amount of the observable for a given treatment.  $\text{Sum}(A)=(A_2+A_1+A_3+A_4)$ ;  $\text{Sum}(B)=(B_1+B_2+B_3+B_4)$ ;  $\text{Sum}(C)=(C_2+C_1+C_3+C_4)$ ;  $\text{Sum}(D)=(D_2+D_1+D_3+D_4)$ ;  $\text{Sum}(E)=(E_2+E_1+E_3+E_4)$ .
2. Sum the sums to obtain the total amount of the observable for all treatments.  $\text{Sum}$ :  $\text{SumT}=\text{Sum}(A)+\text{Sum}(B)+\text{Sum}(C)+\text{Sum}(D)+\text{Sum}(E)$ .
3. Compute the average sum by treatment:  $\text{AvgT}=\text{SumT}/5$ .

4. Compute the differences between the sums on treatment and average sum by treatment:  $Dif(A)=Sum(A)-AvgT$ ;  $Dif(B)=Sum(B)-AvgT$ ;  $Dif(C)=Sum(C)-AvgT$ ;  $Dif(D)=Sum(D)-AvgT$ ;  $Dif(E)=Sum(E)-AvgT$ .

Table 2. Analysis of variance in 5×4 random arrangement experiment

Experimental	A	B	C	D	E	Treatment
Observation 1	3376	3504	3430	3404	3253	
Observation 2	3361	3416	3334	3210	3314	
Observation 3	3366	3244	3291	3168	3287	
Observation 4	3330	3195	3029	3118	3085	
Sum (1 to 4) <sup>(1)</sup> →	13433	13359	13084	12900	12939	<sup>(2)</sup> → $\Sigma = 65715$ ↓ <sup>(3)</sup>
↓ <sup>(5)</sup> $\Sigma(1 \text{ to } 4) - \Sigma/5$	290	216	-59	-243	-204	<sup>(4)</sup> ← $\Sigma/5 = 13143$
$(\Sigma(1 \text{ to } 4) - \Sigma/5)^2$	84100	46656	3481	59049	41616	<sup>(6)</sup> → $\Sigma = 234902$ ↓ <sup>(7)</sup>
↓ <sup>(15)</sup> $SsqE = 289765.75 - 58725.5 = 231040.25$	<sup>(14)</sup> ← $StdD(1..4 \times A..E) = 123.49$					$\Sigma/4 = 58725.5$ ↓ <sup>(8)</sup>
↓ <sup>(16)</sup> $MsqE = SsqE/(19-4) = 15402.68$	$MsqD(1..4 \times A..E) = 15250.83$ ↑ <sup>(13)</sup>					$\Sigma/4/4 = 14681.4$ ↓ <sup>(9)</sup>
↓ <sup>(17)</sup> $StdE = \sqrt{MsqE} = 124.11$	<sup>(11)</sup> $SsqD(1..4 \times A..E) = 289765.75$ ↑ <sup>(12)</sup>					$\sqrt{\Sigma/4/4} = 121.17$
<sup>(10)</sup> degrees of freedom = 19-4	19					4
$F_T = 0.95$ ( $p = 0.4628$ )						

5. Square the differences:  $Dif(A)^2=(Sum(A)-AvgT)^2$ ;  $Dif(B)^2=(Sum(B)-AvgT)^2$ ;  $Dif(C)^2=(Sum(C)-AvgT)^2$ ;  $Dif(D)^2=(Sum(D)-AvgT)^2$ ;  $Dif(E)^2=(Sum(E)-AvgT)^2$ .
6. Sum the difference squared:  $SumDS = Dif(A)^2+Dif(B)^2+Dif(C)^2+Dif(D)^2+Dif(E)^2$ .
7. Compute the sum of squares (SsqT):  $SsqT=SumDS/4$ .
8. Compute the mean of squares (MsqT):  $MsqT=SsqT/4$ .
9. Compute the standard deviation (StdT) due to treatment:  $StdT=\sqrt{MsqT}$ .
10. Compute the degree of freedom (dfT = degree of freedom for treatment; dfEE = degree of freedom for experimental error):  $dfT = \text{number of treatments } (5) - 1$ ;  $dfEE = \text{number of strips } (20) - \text{number of treatments } (5)$ .
11. Compute the sum of squares for the total amount of data:  $SsqD=(A1-SumT/20)^2+(A2-SumT/20)^2+(A3-SumT/20)^2+\dots+(C5-SumT/20)^2$ .
12. Compute the mean of squares for the total amount of data:  $MsqD = SsqD/(dfEE+dfT)$ .
13. Compute the standard deviation for the total amount of data:  $StdD = \sqrt{MsqD}$ .
14. Compute the sum of squares of experimental errors:  $SsqE = SsqD-SsqT$ .
15. Compute the mean of squares of experimental errors:  $MsqE=SsqE/(dfEE-dfT)$ .
16. Compute the standard deviation of the experimental errors:  $StdE=\sqrt{MsqE}$ .

17. Test statistic for treatment:  $F_T = \text{MsqT}/\text{MsqE}$  (follows F distribution with  $\text{dfT}$  and  $(\text{dfT} - \text{dfT} - 1)$  degree of freedom).  $F_T = 0.95$  ( $p = 0.4628$ )

The analysis of the results presented in Table 3 revealed the following classification of errors (Eq. 3.):

$$\text{Error(Treatment)} < \text{Error(Total)} < \text{Error(Experiment)} \quad \text{Eq.3.}$$

The Eq. 3. is the common expected result under assumption that the treatment produces effects on the observable. Thus, an analysis of variance always should produce this relationship in regards of errors.

The analysis of variance for a randomized blocks design is presented in Table 3. The calculations were performed according to the formula used for Table 2.

Table 3. Analysis of Errors: randomized block design

Treatment Block	A	B	C	D	E	SumB	DifB	DifB <sup>2</sup>
1	3314 (A1)	3504 (B1)	3404 (C1)	3029 (D1)	3416 (E1)	16667	238.25	56763.06
2	3361 (A2)	3207 (B2)	3334 (C2)	3287 (D2)	3195 (E2)	16384	-44.75	2002.56
3	3118 (A3)	3167 (B3)	3366 (C3)	3085 (D3)	3244 (E3)	15980	-448.75	201376.56
4	3376 (A4)	3253 (B4)	3291 (C4)	3430 (D4)	3330 (E4)	16680	251.25	63126.56
SumT	13169	13131	13395	12831	13185	65715 (SumTot)		
DifT	26	-12	252	-312	42			
DifT <sup>2</sup>	676	144	63504	97344	1764			
SumDif <sup>2</sup>			163432 (T)	323268.75 (B)				
Ssq			40858 (T)	107756.25 (B)		249607.75 (E)	290465.75 (Tot)	
Msq			10214.50 (T)	35918.75 (B)		20800.65 (E)	15287.67 (Tot)	
Std			101.07 (T)	189.52 (B)		144.22 (E)	123.64 (Tot)	
$F_T = 10214.50 / 20800.65 = 0.49$	$(p = 0.7428)$							
$F_B = 35918.75 / 20800.65 = 1.73$	$(p = 0.2244)$							

T = treatment; B = block number; E = Error; Tot = total amount of variation;

$\text{SsqE} = \text{SsqTot} - \text{SsqT} - \text{SsqB}$ ;  $p = p\text{-value}$ ;  $F = F\text{-test}$ ;

Test statistic for treatment: F dist with  $\text{dfT}$  (degree of freedom for treatment) and (number of observations – number of treatment – number of blocks -1) degree of freedom;

Test statistic for effect of blocks: F dist with  $\text{dfB}$  (degree of freedom for blocks) and (number of observations – number of treatment – number of blocks -1) degree of freedom.

The analysis of the standard deviation resulted on randomized block design allows the following classification of errors (Table 3):

$$\text{Error(Treatment)} < \text{Error(Total)} < \text{Error(Experiment)} < \text{Error(Block)} \quad \text{Eq.4.}$$

The error produce by blocks proved to have the highest value, and it could be explained by the inappropriate analysis of the experimental data or breaking the test assumptions. This kind of analysis is proper when the nuisance factor is known and controllable. If this factor is unknown and uncontrollable, the randomization could balance

their impact across the experiment. The analysis of covariance is preferred if the nuisance factor is known but uncontrollable (the effect of nuisance factor is removed from the analysis). Note that several nuisance factors could combine in a block and influence the analysis of variance.

The test assumption on one-factor randomized block design is violated if [11, 12]: absence of independence between factor and blocking; existence of extreme values (outlier); non-normality of entire sample; patterns in plots of data; small sample size; and/or multiple comparison (more than one primary factor).

### Latin Squares

The results obtained in analysis obtained in investigation of Latin square experimental design are presented in Table 4.

Table 4. Error analysis: Latin square design

	1	2	3	4	5	SumR	DifR	DifR <sup>2</sup>
1	376	371	355	356	335	1793	117.4	13782.76
2	316	338	336	356	332	1678	2.4	5.76
3	326	326	335	343	330	1660	-15.6	243.36
4	317	343	330	327	336	1653	-22.6	510.76
5	321	332	317	318	306	1594	-81.6	6658.56
SumC	1656	1710	1673	1700	1639	8378 (SumTot)		
DifC	-19.6	34.4	-2.6	24.4	-36.6			
DifC <sup>2</sup>	384.16	1183.36	6.76	595.36	1339.56			
SumDif <sup>2</sup>	3509.2 (C)		21201.2 (R)		1651.2 (T)			
Ssq	877.3 (C)		5300.3 (R)		412.8 (T)		436.24 (E) 7026.64 (Tot)	
Msq	219.3 (C)		1325.1 (R)		103.2 (T)		36.4 (E) 292.78 (Tot)	
Std	14.8 (C)		36.4 (R)		10.2 (T)		6.0 (E) 17.1 (Tot)	
F <sub>T</sub>	= 103.2 / 36.4 = 2.84 (p = 0.0618)							

C = column; R = row; T = treatment; E = error; Tot = total

The analysis on Latin square design allows the following classification of errors (Table 4):

$$\text{Error(Experiment)} < \text{Error(Treatment)} < \text{Error(Total)} < \text{Error(Column)} < \text{Error(Row)} \quad \text{Eq.5.}$$

The analysis of error classification according with the investigate designs of experiments revealed the following:

- The randomized arrangements and randomized block design produce the same classification of errors.

- The smallest error proved to be of treatment in randomized arrangement and randomized block design and of experiment in Latin square.
- The error of experiment proved to have the smaller value in Latin square design but the highest value in randomized blocks and randomized arrangements.

There are a lot of programs which allows the analysis of error and variance according to not to experimental design [13-18]. A series of designs also were developed and applied in order to reduce different kind of errors and to respond to different restrictions: independent replications and the same square used in replications [19]; semi-Latin squares [20, 21]; split plot experimental design [22]; clustered versus spatial randomized blocks [23]; non-randomized block design [24]. New statistical techniques were proposed to be used for analysis of data resulted from randomized block designs and Latin squares: Sign and Hotelling-Hsu's  $T^2$  tests [25]; Friedman and Page tests [26]; rank test for a randomized block design based on the quantile score used when a small proportion of the observations from each treatment group shows very high response values or when the response is dichotomous [27]; locally most powerful (LMP) rank test [28]. In these situations, the researchers must have wide knowledge in at least 2 domains in order to be able to conduct a correct experiment: design of experiment and statistical analysis.

The field experiments should be correctly designed and analyzed in order to provide the maximum amount of information for the minimum amount of resources. The main criteria for a well-designed experiment recommended to be carefully addressed are:

1. Unbiased (identical as possible environmental conditions): allows an accurate and valid comparison between treatment groups.
2. High precision (uniform selection of experimental material; increasing the number of observations): allows identifying the true effect if any exists.
3. Wide target of applicability: allows exploration of other nuisance factors.
4. Simple experiments (as simple as possible).

Last but not the least, a well-design experiment should provide all data necessary to calculate uncertainty (capable to be statistically analyzed and to quantify the level of confidence in the results) without which the experiment is useless.

## **Conclusions**

A correct experimental design is as important as a correct statistical analysis in order to obtain valid and reliable conclusion from field experiments. Certain restrictions must be imposed when the plots are arranged in order to be able to accurately estimate the errors.

The Latin square design proved to have the smallest experimental error compared to randomized arrangement and randomized block design. The classification of errors proved to be similar in randomized arrangements and randomized block design.

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